

SPECIAL ARTICLE

Explaining the Decrease in U.S. Deaths from Coronary Disease, 1980–2000

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ABSTRACT

BACKGROUND

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Mortality from coronary heart disease in the United States has decreased substantially in recent decades. We conducted a study to determine how much of this decrease could be explained by the use of medical and surgical treatments as opposed to changes in cardiovascular risk factors.

METHODS

We applied a previously validated statistical model, IMPACT, to data on the use and effectiveness of specific cardiac treatments and on changes in risk factors between 1980 and 2000 among U.S. adults 25 to 84 years old. The difference between the observed and expected number of deaths from coronary heart disease in 2000 was distributed among the treatments and risk factors included in the analyses.

RESULTS

From 1980 through 2000, the age-adjusted death rate for coronary heart disease fell from 542.9 to 266.8 deaths per 100,000 population among men and from 263.3 to 134.4 deaths per 100,000 population among women, resulting in 341,745 fewer deaths from coronary heart disease in 2000. Approximately 47% of this decrease was attributed to treatments, including secondary preventive therapies after myocardial infarction or revascularization (11%), initial treatments for acute myocardial infarction or unstable angina (10%), treatments for heart failure (9%), revascularization for chronic angina (5%), and other therapies (12%). Approximately 44% was attributed to changes in risk factors, including reductions in total cholesterol (24%), systolic blood pressure (20%), smoking prevalence (12%), and physical inactivity (5%), although these reductions were partially offset by increases in the body-mass index and the prevalence of diabetes, which accounted for an increased number of deaths (8% and 10%, respectively).

CONCLUSIONS

Approximately half the decline in U.S. deaths from coronary heart disease from 1980 through 2000 may be attributable to reductions in major risk factors and approximately half to evidence-based medical therapies.

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RATES OF DEATH FROM CORONARY HEART disease in the United States underwent profound secular changes during the 20th century.^{1,2} After peaking around 1968, age-adjusted rates were cut in half. Two factors may have contributed to this decline.

First, there have been substantial decreases in the prevalence of some major cardiovascular risk factors, including smoking, elevated total cholesterol, and high blood pressure.³⁻⁸ However, the prevalence of both obesity and diabetes has increased alarmingly.⁹⁻¹¹

Second, there has been a revolution in the treatments for established coronary heart disease, with major breakthroughs in evidence-based therapies, including the use of thrombolysis, coronary-artery bypass grafting (CABG), coronary angioplasty and stents, and angiotensin-converting-enzyme (ACE) inhibitors and statins.

The annual direct and indirect costs for coronary heart disease were \$142.5 billion in 2006, and they continue to rise.¹² Determining the respective contributions of prevention and therapy to the declines in mortality from coronary heart disease is therefore becoming increasingly important, for the purposes of both understanding past trends and planning future strategies. Estimates of the contribution from reductions in risk factors before 1990 have ranged from 50 to 54% in the United States^{13,14} and from 44 to 76% in other industrialized countries.¹⁵⁻²² However, to our knowledge, no U.S. studies have considered the dramatic changes since 1990 or have attempted to quantify the relative contributions of specific therapies and trends in risk factors. We therefore applied a model that has been used successfully in several other countries to examine trends in U.S. deaths from coronary heart disease between 1980 and 2000.

METHODS

MORTALITY MODEL AND DATA SOURCES

To examine the contributions of various factors to the changes in rates of death from coronary heart disease among U.S. adults 25 to 84 years of age, we used an updated version of the IMPACT mortality model, which was previously validated in Europe, New Zealand, and China.¹⁸⁻²³ This model has been described in detail elsewhere.^{18,19,23,24} It incorporates major population risk factors for coronary heart disease (smoking, high blood pres-

sure, elevated total cholesterol, obesity, diabetes, and physical inactivity) and all usual medical and surgical treatments for coronary heart disease.

Wherever possible, data sources specific to the U.S. population were used to construct the U.S. model. When more than one data source was available, we chose the source that we considered to be most representative, least biased, and most up-to-date. Detailed information on the IMPACT model and data sources for the U.S. analysis is provided in the Supplementary Appendix, available with the full text of this article at www.nejm.org.

DEATHS PREVENTED OR POSTPONED

Data on the total U.S. population and age distribution in 1980 and 2000 were obtained from the U.S. Census Bureau. Deaths according to age and sex and mortality rates associated with coronary heart disease in 1980 and 2000 were obtained from the National Vital Statistics System of the National Center for Health Statistics. We calculated the number of deaths from coronary heart disease that would have been expected in 2000 if the mortality rates in 1980 had remained unchanged by multiplying the age-specific mortality rates for 1980 by the population for each 10-year age stratum in the year 2000 (thus accounting for the aging of the population). Subtracting the number of deaths observed in 2000 from the number expected then yielded the drop in the number of deaths (prevented or postponed) in 2000 that the model would have to explain.

TREATMENTS AND MORTALITY REDUCTIONS

The prevalence of coronary heart disease by diagnosis, the estimated frequency of use of specific treatments, the case fatality rate by diagnosis, and the risk reduction due to treatment, all stratified by age and sex, were obtained from published sources (Tables 2 through 5 in the Supplementary Appendix). The number of deaths prevented or postponed as a result of each intervention in each group of patients in the year 2000 (Table 1) was calculated by multiplying the number of people in each diagnostic group by the proportion of those patients who received a particular treatment, by the case fatality rate over a period of 1 year, and by the relative reduction in the 1-year case fatality rate that was accounted for by the treatment.^{19,20} For example, in the United States in 2000, approximately 102,280 men between the

Table 1. Estimated Deaths Prevented or Postponed by Medical or Surgical Treatments in the United States in 2000.*

Treatment	No. of Eligible Patients	Patients Receiving Treatment	Relative Risk Reduction	Mean Case Fatality Rate	Absolute Risk Reduction	Deaths Prevented or Postponed			percent of total reduction		
						Best Estimate	Minimum Estimate	Maximum Estimate	Best Estimate	Minimum Estimate	Maximum Estimate
Acute myocardial infarction	670,715	—	—	0.094	—	21,570	9,045	37,720	6.3	2.6	11.0
Resuscitation in the community	204,330	43	0.05	0.094	0.050	4,435	2,205	7,450	1.3	0.6	2.2
Resuscitation in the hospital	13,415	100	0.33	0.094	0.342	4,095	1,050	5,660	1.2	0.3	1.7
Thrombolysis	670,715	20	0.24	0.094	0.019	2,410	1,260	4,250	0.7	0.4	1.2
Aspirin	670,715	84	0.15	0.094	0.014	7,735	3,960	13,265	2.3	1.2	3.9
Beta-blocker	670,715	7	0.31	0.094	0.025	340	-505	2,130	0.3	-0.1	0.6
ACE inhibitor	670,715	16	0.04	0.094	0.004	805	160	1,995	0.1	0.0	0.6
Primary angioplasty	670,715	21	0.07	0.094	0.007	1,070	560	1,900	0.2	0.2	0.6
Primary CABG	670,715	8	0.39	0.094	0.037	1,925	1,035	3,495	0.6	0.3	1.0
Treatments in 1980 subtracted						-1,245	-680	-2,415	-0.4	-0.2	-0.7
Unstable angina	665,260	60	0.33	0.065	0.021	13,575	6,960	23,490	4.0	2.0	6.9
Aspirin and heparin		17	0.15	0.065	0.010	6,350	3,250	10,975	1.9	1.0	3.2
Aspirin alone		21	0.09	0.065	0.006	790	410	1,385	0.2	0.1	0.4
Glycoprotein IIb/IIIa antagonists and clopidogrel		21	0.09	0.065	0.006	595	310	1,040	0.2	0.1	0.3
CABG		20	0.43	0.065	0.028	2,760	1,415	4,765	0.8	0.4	1.4
Angioplasty		30	0.32	0.065	0.021	3,080	1,575	5,320	0.9	0.5	1.6
Secondary prevention after myocardial infarction	2,866,965	28,565	0.057	0.057		28,565	12,255	62,030	8.4	3.6	18.2
Aspirin		38	0.15	0.057	0.007	5,135	2,285	11,560	1.5	0.7	3.4
Beta-blocker		29	0.23	0.057	0.012	6,750	2,765	13,995	2.0	0.8	4.1
ACE inhibitor		26	0.20	0.057	0.010	5,155	2,125	10,745	1.5	0.6	3.1
Statin		36	0.22	0.057	0.011	4,700	2,175	11,000	1.4	0.6	3.2
Warfarin		9	0.22	0.057	0.012	2,175	870	4,395	0.6	0.3	1.3
Rehabilitation		21	0.26	0.057	0.012	4,650	2,040	10,330	1.4	0.6	3.0

Secondary prevention after CABG or PTCA	1,948,660	0.019	7,435	3,070	15,535	2.2	0.9	4.6
Aspirin	40	0.15	1,310	555	2,800	0.4	0.2	0.8
Beta-blocker	33	0.23	1,460	595	3,010	0.4	0.2	0.9
ACE inhibitor	26	0.20	1,010	410	2,080	0.3	0.1	0.6
Statin	38	0.22	1,550	665	3,365	0.5	0.2	1.0
Warfarin	9	0.22	450	180	915	0.1	0.1	0.3
Rehabilitation	32	0.26	1,655	665	3,360	0.5	0.2	1.0
Secondary-prevention treatments in 1980 subtracted			-195					
Chronic angina		0.015	17,730	5,595	40,950	5.2	1.6	12.0
CABG, 1990 to 2000	4,987,375	0.020	14,365	6,720	34,180	4.2	2.0	10.0
With CABG in 1980 subtracted	2,356,695		-3,065	-1,960	-4,410	-0.9	-0.6	-1.3
Angioplasty, 1990 to 2000	22,059,760	0.019	4,390	0	6,960	1.3	0.0	2.0
Aspirin in the community	2,119,495	0.011	1,060	435	2,195	0.3	0.1	0.6
Statins in the community	2,119,495	0.011	980	400	2,030	0.3	0.1	0.6
Heart failure with hospital admission		0.237	11,735	4,945	25,030	3.4	1.4	7.3
ACE inhibitor	258,745	0.20	3,130	1,375	6,960	0.9	0.4	2.0
Beta-blocker	43	0.35	4,300	1,750	8,850	1.3	0.5	2.6
Spironolactone	7	0.30	825	330	1,665	0.2	0.1	0.5
Aspirin	30	0.15	1,700	715	3,615	0.5	0.2	1.1
Statins	28	0.23	1,785	780	3,940	0.5	0.2	1.2
Heart failure in the community	1,876,405	0.063	18,500	7,735	39,170	5.4	2.3	11.5
ACE inhibitor	48	0.20	3,375	1,550	7,845	1.0	0.5	2.3
Beta-blocker	30	0.35	7,180	2,940	14,885	2.1	0.9	4.4
Spironolactone	8	0.28	1,685	690	3,500	0.5	0.2	1.0
Aspirin	30	0.15	3,105	1,265	6,405	0.9	0.4	1.9
Statin	30	0.23	3,150	1,290	6,535	0.9	0.4	1.9
Hypertension	54,353,660	0.13	23,845	1,945	68,370	7.0	0.6	20.0
Statins for lipid reduction (primary prevention)†	96,384,630	0.29	16,580	6,515	35,100	4.9	1.9	10.3
Total treatments			159,330	58,065	347,395	46.6	19.2	94.3

* Percentages may not sum to 100 because of rounding. Data sources are described in the Supplementary Appendix. CABG denotes coronary-artery bypass grafting, AMI acute myocardial infarction, ACE angiotensin-converting enzyme, and PTCA percutaneous transluminal coronary angioplasty (with or without stenting).

† The number of deaths prevented or postponed includes 475 that were prevented or postponed owing to treatment with gemfibrozil and niacin for primary prevention of hyperlipidemia.

ages of 55 and 64 years were hospitalized with acute myocardial infarction. Some 84% were given aspirin, with an expected mortality reduction of 15%.²⁵ The expected age-specific, 1-year case fatality rate was approximately 5.4%.²⁶ The number of deaths prevented or postponed for at least a year by the use of aspirin among men in this age group was then calculated as $102,280 \times 0.84 \times 0.15 \times 0.054 = 696$.

Several adjustments were made to these basic analyses. Although most of the therapeutic measures studied were not in use in 1980, in some cases such use was already substantial (e.g., CABG for stable angina pectoris). In such cases, the number of deaths prevented or postponed as a result of the therapy as used in 1980 was calculated and subtracted from the number of deaths for 2000 to calculate the net benefit. We assumed that compliance — the proportion of treated patients actually taking therapeutically effective levels of medication — was 100% among hospitalized patients, 70% among symptomatic patients in the community, and 50% among asymptomatic patients in the community.^{19,24,27,28} To avoid double counting of patients treated, we identified potential overlaps between different groups of patients and made appropriate adjustments (Table 9 in the Supplementary Appendix). For example, heart failure develops within 1 year after acute myocardial infarction in approximately one quarter of survivors, and approximately half the patients undergoing CABG have had a previous myocardial infarction.^{19,24} To address the potential effect on the relative reduction in the case fatality rate for individual patients receiving multiple treatments, we used the Mant and Hicks cumulative-relative-benefit approach²⁹:

$$\text{relative benefit} = 1 - \frac{(1 - \text{relative reduction in case fatality rate for treatment A}) \times (1 - \text{relative reduction in case fatality rate for treatment B}) \times \dots \times (1 - \text{relative reduction in case fatality rate for treatment N})}{1 - \text{relative reduction in case fatality rate for treatment A}}$$

RISK FACTORS AND MORTALITY REDUCTIONS

Two approaches were used to calculate the numbers of deaths prevented or postponed as a result of changes in risk factors. We used a regression approach for systolic blood pressure, cholesterol, and body-mass index. The number of deaths prevented or postponed as a result of the change in the prevalence of or mean value for each of these risk factors (Table 2) was estimated as the prod-

uct of three variables: the number of deaths from coronary heart disease in 1980 (the base year), the subsequent reduction in that risk factor (Table 2 in the Supplementary Appendix), and the regression coefficient quantifying the change in mortality from coronary heart disease per unit of absolute change in the risk factor (Table 6 in the Supplementary Appendix). For example, in 1980, there were 26,352 deaths from coronary heart disease among 12,629,000 women who were 55 to 64 years of age. The mean systolic blood pressure in this group decreased by 3.09 mm Hg between 1980 and 2000. The largest meta-analysis showed an estimated age- and sex-specific reduction in mortality of 50% for every reduction of 20 mm Hg in systolic pressure, yielding a logarithmic (ln) coefficient of -0.035 .³³

The number of deaths prevented or postponed as a result of this change was then estimated as follows:

$$\begin{aligned} \text{number of deaths} &= (1 - e^{(\text{coefficient} \times \text{change})}) \times \text{deaths in 1980} \\ &= (1 - e^{(-0.035 \times 3.09)}) \times 26,352 = 2701. \end{aligned}$$

The population-attributable risk fraction was used to determine the effect of changes in the prevalence of smoking, diabetes, and physical inactivity. The population-attributable risk fraction was calculated conventionally as $[P \times (RR-1)] \div [(1+P) \times (RR-1)]$, where P is the prevalence of the risk factor (Table 2 in the Supplementary Appendix) and RR is the relative risk of death from coronary heart disease associated with that risk factor (Table 7 in the Supplementary Appendix). The number of deaths prevented or postponed was then estimated as the number of deaths from coronary heart disease in 1980 (the base year) multiplied by the difference between the population-attributable risk fraction in 1980 and that in 2000 (Table 2). For example, the prevalence of diabetes among men 65 to 74 years of age increased from 14.5% in 1980 to 20.7% in 2000. Given a relative risk of 1.93, the population-attributable risk fraction increased from 0.119 to 0.161. Additional deaths from coronary heart disease in 2000 that were attributable to an increased prevalence of diabetes were therefore calculated as follows^{18,19,23,24}:

$$\begin{aligned} \text{deaths from coronary heart disease in 1980} &= (123,055) \times \\ & (0.161 - 0.119) = 5168. \end{aligned}$$

Because independent regression coefficients and relative risks for each risk factor were obtained from multivariate analyses, we assumed

Table 2. Deaths from Coronary Heart Disease That Were Prevented or Postponed as a Result of Changes in Population Risk Factors in the United States, 1980 to 2000.*

Risk Factor†	Absolute Level of Risk Factor‡		Change in Risk Factor		Beta Regression Coefficient for Change in Mortality Rate§	Relative Risk		Deaths Prevented or Postponed			
	1980	2000	Absolute Change	Relative Change (%)		Best Estimate	Minimum Estimate	Maximum Estimate	Best Estimate	Minimum Estimate	Maximum Estimate
Smoking prevalence (%)	36.3	24.6	-11.7	-32.2		39,925	34,955	52,435	11.7	10.2	15.3
Men					2.52						
Women					2.14						
Systolic blood pressure (mm Hg)	129.0	123.9	-5.1	-4.0		68,800	53,730	105,060	20.1	15.7	30.7
Men					-0.0334						
Women					-0.0413						
Total cholesterol (mmol/liter)	5.67	5.33	-0.34	-6.1		82,830	58,455	95,570	24.2	17.1	28.0
Men					-0.9458						
Women					-0.9121						
Physical inactivity (%)	29.6	27.3	-2.3	-7.8		17,445	8,340	29,035	5.1	2.4	8.5
Men					1.27						
Women					1.33						
BMI	25.6	28.2	+2.6	10.1		-25,905	-14,430	-40,405	-7.6	-4.2	-11.8
Men					0.0297						
Women					0.0297						
Diabetes prevalence (%)	6.5	9.4	+2.9	44.2		-33,465	-23,885	-43,330	-9.8	-7.0	-12.7
Men					1.93						
Women					2.59						
Total risk factors						149,635	117,165	198,360	43.8	34.3	58.0

* Percentages may not sum to 100 because of rounding. BMI denotes body-mass index (the weight in kilograms divided by the square of the height in meters). To convert the values for cholesterol to milligrams per deciliter, divide by 0.02586. Data sources are described in the Supplementary Appendix.

† The total adult population in 1980 was 177,745,055. For systolic pressure, the numbers of deaths exclude patients receiving treatment for hypertension, and for total cholesterol, the numbers exclude patients receiving statins.

‡ Data are from the National Center for Health Statistics,^{30,31} except for data on physical inactivity, which are from the Behavioral Risk Factor Surveillance System.³²

§ The change in the mortality rate per unit of measurement for the risk factor is shown.

that there was no further synergy between the treatment and risk-factor sections of the model or among the major risk factors.

The number of deaths prevented or postponed as a result of changes in risk factors was systematically quantified for each specific patient group to account for potential differences in effect. Lag times between the change in the risk-factor rate and the change in the event rate were not modeled; it was assumed that these lag times would be relatively unimportant over a period of two decades.^{20,23,34,35}

COMPARISON OF ESTIMATED AND OBSERVED MORTALITY CHANGES

The model estimates for the total number of deaths prevented or postponed by each treatment and for each risk-factor change were rounded to the nearest multiple of 5 (e.g., 696 became 695). All these figures were then summed and compared with the observed changes in mortality for men and women in each age group. Any shortfall in the overall model estimate was then presumed to be attributable either to inaccuracies in our calculated estimates or to other, unmeasured risk factors.^{19,20,24}

SENSITIVITY ANALYSES

We tested all the above assumptions and variables in a multiple-way sensitivity analysis, using the analysis-of-extremes method.^{19,20,24,36} For each variable in the model, we assigned a lower value

and an upper value, using 95% confidence intervals when available and otherwise using $\pm 20\%$ (for the number of patients, use of treatment, and compliance). For example, for aspirin treatment in men 55 to 64 years of age who were hospitalized with acute myocardial infarction, the best estimate was 696 deaths prevented or postponed. The minimum estimate from the multiple-way sensitivity analysis was 259, and the maximum estimate was 1501 (Table 3).

RESULTS

From 1980 to 2000, the age-adjusted rate of coronary heart disease fell from 542.9 to 266.8 cases per 100,000 population among men aged 25 to 84 years and from 263.3 to 134.4 among women aged 25 to 84 years. In 1980, a total of 462,984 deaths among people in this age group were recorded as due to coronary heart disease, according to the *International Classification of Diseases, 9th Revision* (codes 410–414 and 429.2).⁴¹ In 2000, a total of 337,658 such deaths were recorded, according to the *International Classification of Diseases, 10th Revision* (codes I20–I25).⁴² However, had the age-specific death rates from 1980 remained in 2000, an additional 341,745 deaths from coronary heart disease would have occurred.

The U.S. IMPACT model explained approximately 308,965 (90%) of this decrease in the number of deaths from coronary heart disease. Under the assumptions of the sensitivity analysis,

Table 3. Example of a Multiple-Way Sensitivity Analysis.*

Estimate	No. of Patients (a) [†]	Proportion Receiving Treatment (b) [‡]	Relative Mortality Reduction (c) [‡]	1-Year Case Fatality Rate (d) [‡]	No. of Deaths Prevented or Postponed (a×b×c×d)
			percent		
Best	102,280	0.84	15	5.4	696
Minimum	81,824	0.67	11	4.3	259
Maximum	122,736	0.99	19	6.5	1501

* In the United States in 2000, about 102,280 men aged 55 to 64 years were hospitalized with acute myocardial infarction, of whom approximately 84% were given aspirin. Aspirin use reduced the case fatality rate by approximately 15%. The underlying 1-year case fatality rate in these men was approximately 5.4%. The calculated number of deaths prevented or postponed was approximately 696. A multiple-way sensitivity analysis was then performed. Lower and upper bounds for each variable were estimated with use of 95% confidence intervals, when available, or failing that, with use of calculated bounds of $\pm 20\%$ (treatment uptake, however, was capped at 99%). Multiplying all lower-bound estimates together yielded the lower-bound estimate of deaths prevented or postponed, and multiplying all upper-bound estimates together yielded the upper-bound estimate of deaths prevented or postponed.

[†] Numbers of patients are from the National Hospital Discharge Survey³⁷ and the Medical Expenditure Panel Survey.³⁸
[‡] Treatment data are from Rogers et al.,³⁹ data on mortality reduction are from the Antithrombotic Trialists' Collaboration,⁴⁰ and case fatality rates are from Capewell et al.²⁶

the minimum and maximum numbers of deaths from coronary heart disease that were explained were 175,230 (51%) and 545,755 (160%). The agreement between the number of estimated deaths and the number of observed deaths was reasonably good for men across all groups and for women under the age of 75 years (Fig. 1). Changes in medical treatments accounted for approximately 47% and risk-factor changes accounted for approximately 44% of the decrease in deaths (Tables 1 and 2).

MEDICAL AND SURGICAL TREATMENTS

Approximately 159,330 of the deaths from coronary heart disease that were prevented or postponed were attributable to medical therapies (minimum estimate, 58,065; maximum estimate, 347,395) (Table 1). The largest reductions in deaths came from the use of secondary-prevention medications or rehabilitation after acute myocardial infarction or after revascularization (a total reduction of approximately 35,800 deaths) and from the use of initial treatments for acute myocardial infarction or unstable angina (approximately 35,145 deaths), followed by treatments for heart failure and hypertension, statin therapy for primary prevention, and treatments for chronic angina. The use of revascularization for chronic angina resulted in a reduction of approximately 15,690 deaths in 2000, as compared with deaths in 1980, or approximately 5% of the total.

RISK FACTORS

Approximately 149,635 fewer deaths from coronary heart disease were attributable to changes in risk factors (minimum estimate, 117,165; maximum estimate, 198,360) (Table 2). Decreases in the total cholesterol concentration (by 0.34 mmol per liter), systolic blood pressure (by 5.1 mm Hg), and smoking prevalence (by 11.7%) were estimated to have prevented or postponed approximately 82,830, 68,800, and 39,925 deaths, respectively. The 2.3% decrease in physical inactivity prevented or postponed approximately 17,445 deaths. In contrast, the increase in the body-mass index (the weight in kilograms divided by the square of the height in meters) of 2.6 and the 2.9% increase in the prevalence of diabetes resulted in approximately 25,905 and 33,465 additional deaths overall, respectively (Table 2).

PROPORTIONAL CONTRIBUTIONS TO THE DECREASE IN DEATHS

Sensitivity analyses showed that the proportional contributions of specific treatments and risk-factor changes to the overall reduction in deaths from coronary heart disease in 2000 were relatively consistent (Tables 1 and 2). Thus, all initial treatments for acute myocardial infarction together accounted for approximately 21,570 fewer deaths, representing 6.3% of the total decrease of 341,745 deaths. The minimum estimated contribution was 9045 fewer deaths (2.6%), and the maximum was 37,720 (11.0%). The contribution of treatments for acute myocardial infarction therefore remained consistently smaller than that of secondary prevention or therapies for heart failure, irrespective of whether best, minimum, or maximum estimates were compared (Table 1).

DISCUSSION

The burden of coronary heart disease in the United States remains enormous, even though associated mortality rates fell by more than 40% between 1980 and 2000. These two decades saw rapid growth in costly medical technology and pharmaceutical treatments for coronary heart disease, as well as substantial public health efforts to reduce

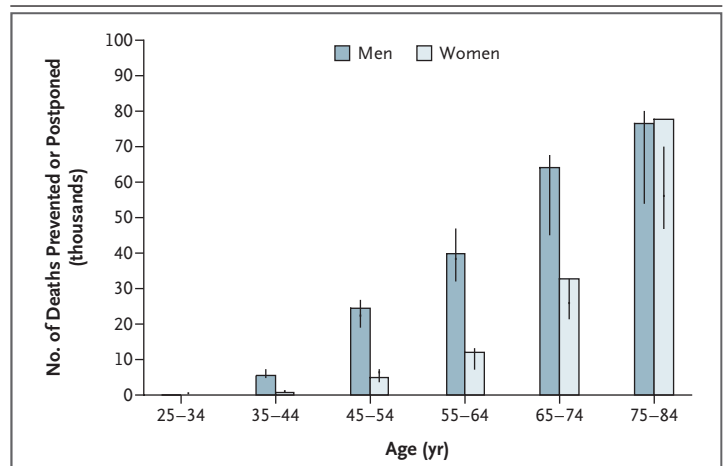


Figure 1. Estimated and Observed Reductions in Deaths from Coronary Heart Disease in the United States between 1980 and 2000, Stratified According to Age and Sex.

The bars show the observed decrease in deaths in each age group, and the vertical lines the extreme minimum and maximum estimates in the sensitivity analysis.

the prevalence of major cardiovascular risk factors. Establishing the relative contributions of these two approaches is therefore of considerable importance. We found that reductions in major risk factors probably accounted for approximately half the decrease in deaths from coronary heart disease, as in most other industrialized countries studied.¹⁵⁻²² Earlier U.S. studies likewise suggested a contribution of approximately 54% of the reduction in deaths between 1968 and 1976¹⁴ and approximately 50% between 1980 and 1990.¹³

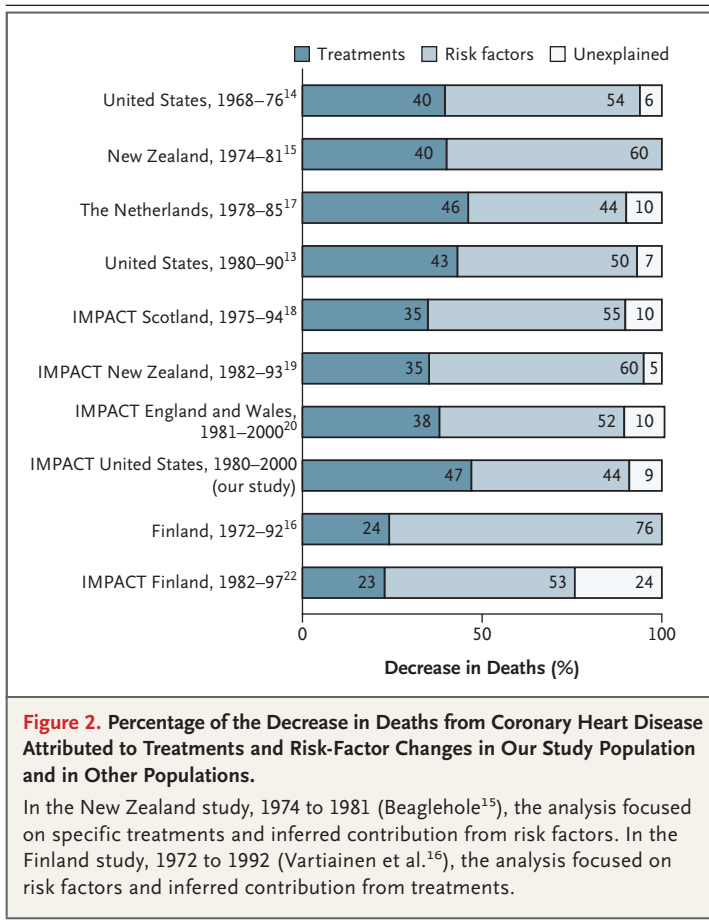
Irrespective of the assumptions used, we found that the largest contributions from medical therapies consistently came from secondary prevention, followed by treatments for acute coronary syndromes, then heart failure. Revascularization by means of CABG or angioplasty for stable or unstable disease together accounted for approximately 7% of the overall drop in deaths from coronary heart disease, a finding that is consistent with the results of previous studies in the United States⁴³ and elsewhere.^{19-22,44}

Although most of the changes in treatments and risk factors between 1980 and 2000 led to reductions in deaths from coronary heart disease, two major exceptions are noteworthy. Our analysis estimated that increases in the body-mass index accounted overall for about 26,000 additional deaths from coronary heart disease in 2000 and increases in the prevalence of diabetes for about 33,500 additional deaths; both figures are consistent with the results of other recent studies.^{45,46} Efforts to address these two risk factors should therefore receive particular attention in future measures to improve the public health.^{10,11}

Modeling studies have a number of potential strengths, including the ability to transparently integrate and simultaneously consider huge amounts of data from many sources and then test explicit assumptions by means of sensitivity analyses. Our analysis of extremes suggested that the proportional contributions to the overall reductions in deaths from specific treatments and risk-factor changes remained reasonably consistent, irrespective of whether best, minimum, or maximum estimates were considered (Tables 1 and 2). This was reassuring, as was the general consistency with the results of most studies performed elsewhere (Fig. 2).^{15-17,19,20}

However, all modeling analyses should be interpreted with appropriate caution. All require the gathering of data from numerous sources, each with recognized limitations. We sometimes had to use data from studies that might have been limited by geographic, ethnic, or selection bias or by the need to extrapolate to older age groups. Risk estimates were not necessarily fully independent of each other. Furthermore, most interactions were averaged across broad groups. We therefore made the explicit assumptions detailed in the Supplementary Appendix. Furthermore, we analyzed only the estimated reduction in deaths from coronary heart disease, not life-years gained or improvement in the quality of life.⁴⁷ Analyses of these changes are warranted, as well as comparisons among racial and ethnic groups and economic analyses.

The estimates of changes in risk factors remain imprecise. Furthermore, we did not explicitly consider the effect of lag times; however, they may be relatively unimportant over a 20-year period.^{20,23,33,35} Although major efforts were made to address overlaps, residual double counting of some individual patients remains possible. We



also assumed that, after adjustments for reduced dosing and imperfect compliance, the efficacy of treatments in randomized, controlled trials could be generalized to usual clinical practice.^{48,49} Both assumptions may have potentially overestimated the true treatment effect.

In conclusion, our analyses suggest that approximately half the recent decrease in deaths from coronary heart disease in the United States may be attributable to reductions in major risk factors and approximately half to evidence-based

medical therapies. Future strategies for preventing and treating coronary heart disease should therefore be comprehensive, maximizing the coverage of effective treatments and actively promoting population-based prevention by reducing risk factors.

No potential conflict of interest relevant to this article was reported.

The findings and conclusions in this article are those of the authors and do not represent the views of the Centers for Disease Control and Prevention.

REFERENCES

- Rosamond W, Flegal K, Friday G, et al. Heart disease and stroke statistics — 2007 update: a report from the American Heart Association Statistics Committee and Stroke Statistics Subcommittee. *Circulation* 2007;115:e69-e171. [Erratum, *Circulation* 2007;115:e172].
- Morbidity and mortality: 2004 chart book on cardiovascular, lung, and blood diseases. Bethesda, MD: National Heart, Lung, and Blood Institute, 2004. (Accessed May 11, 2007, at <http://www.nhlbi.nih.gov/resources/docs/cht-book.htm>.)
- Centers for Disease Control and Prevention. Percentage of adults who were current, former, or never smokers, overall and by sex, race, Hispanic origin, age, and education status: National Health Interview Surveys, selected years — United States, 1965–2004. (Accessed May 11, 2007, at http://www.cdc.gov/tobacco/data_statistics/tables/adult/table_2.htm.)
- Hajjar I, Kotchen TA. Trends in prevalence, awareness, treatment, and control of hypertension in the United States, 1988–2000. *JAMA* 2003;290:199–206.
- Burt VL, Cutler JA, Higgins M, et al. Trends in the prevalence, awareness, treatment, and control of hypertension in the adult US population: data from the Health Examination Surveys, 1960 to 1991. *Hypertension* 1995;26:60–9. [Erratum, *Hypertension* 1996;27:1192.]
- Johnson CL, Rifkind BM, Sempos CT, et al. Declining serum total cholesterol levels among US adults: the National Health and Nutrition Examination Surveys. *JAMA* 1993;269:3002–8.
- Ford ES, Mokdad AH, Giles WH, Mensah GA. Serum total cholesterol concentrations and awareness, treatment, and control of hypercholesterolemia among US adults: findings from the National Health and Nutrition Examination Survey, 1999 to 2000. *Circulation* 2003;107:2185–9.
- Prevalence of no leisure-time physical activity — 35 states and the District of Columbia, 1988–2002. *MMWR Morb Mortal Wkly Rep* 2004;53:82–6.
- Hedley AA, Ogden CL, Johnson CL, Carroll MD, Curtin LR, Flegal KM. Prevalence of overweight and obesity among US children, adolescents, and adults, 1999–2002. *JAMA* 2004;291:2847–50.
- Harris MI, Hadden WC, Knowler WC, Bennett PH. Prevalence of diabetes and impaired glucose tolerance and plasma glucose levels in U.S. population aged 20–74 yr. *Diabetes* 1987;36:523–34.
- Harris MI, Flegal KM, Cowie CC, et al. Prevalence of diabetes, impaired fasting glucose, and impaired glucose tolerance in U.S. adults: the Third National Health and Nutrition Examination Survey, 1988–1994. *Diabetes Care* 1998;21:518–24.
- Thom T, Haase N, Rosamond W, et al. Heart disease and stroke statistics — 2006 update: a report from the American Heart Association Statistics Committee and Stroke Statistics Subcommittee. *Circulation* 2006;113:e696, 114:e630.]
- Hunink MG, Goldman L, Tosteson AN, et al. The recent decline in mortality from coronary heart disease, 1980–1990: the effect of secular trends in risk factors and treatment. *JAMA* 1997;277:535–42.
- Goldman L, Cook EF. The decline in ischemic heart disease mortality rates: an analysis of the comparative effects of medical interventions and changes in lifestyle. *Ann Intern Med* 1984;101:825–36.
- Beaglehole R. Medical management and the decline in mortality from coronary heart disease. *Br Med J (Clin Res Ed)* 1986;292:33–5.
- Vartiainen E, Puska P, Pekkanen J, Tuomilehto J, Jousilahti P. Changes in risk factors explain changes in mortality from ischaemic heart disease in Finland. *BMJ* 1994;309:23–7.
- Bots ML, Grobbee DE. Decline of coronary heart disease mortality in the Netherlands from 1978 to 1985: contribution of medical care and changes over time in presence of major cardiovascular risk factors. *J Cardiovasc Risk* 1996;3:271–6.
- Capewell S, Morrison CE, McMurray JJ. Contribution of modern cardiovascular treatment and risk factor changes to the decline in coronary heart disease mortality in Scotland between 1975 and 1994. *Heart* 1999;81:380–6.
- Capewell S, Beaglehole R, Seddon M, McMurray J. Explanation for the decline in coronary heart disease mortality rates in Auckland, New Zealand, between 1982 and 1993. *Circulation* 2000;102:1511–6.
- Unal B, Critchley JA, Capewell S. Explaining the decline in coronary heart disease mortality in England and Wales, 1981–2000. *Circulation* 2004;109:1101–7.
- Idem*. Modelling the decline in coronary heart disease deaths in England and Wales, 1981–2000: comparing contributions from primary prevention and secondary prevention. *BMJ* 2005;331:614.
- Laatikainen T, Critchley J, Vartiainen E, Salomaa V, Ketonen M, Capewell S. Explaining the decline in coronary heart disease mortality in Finland between 1982 and 1997. *Am J Epidemiol* 2005;162:764–73.
- Critchley J, Liu J, Zhao D, Wei W, Capewell S. Explaining the increase in coronary heart disease mortality in Beijing between 1984 and 1999. *Circulation* 2004;110:1236–44.
- Ural B, Critchley J, Capewell S. IMPACT, a validated, comprehensive coronary heart disease model. Liverpool, United Kingdom: University of Liverpool, 2006. (Accessed May 11, 2007, at <http://www.liv.ac.uk/PublicHealth/sc/bua/impact.html>.)
- Ryan R, Majeed A. Prevalence of ischaemic heart disease and its management with statins and aspirin in general practice in England and Wales, 1994–98. *Health Stat Q* 2001;12:34–9.
- Capewell S, Livingston BM, MacIntyre K, et al. Trends in case-fatality in 117 718 patients admitted with acute myocardial infarction in Scotland. *Eur Heart J* 2000;21:1833–40.
- Butler J, Arbogast PG, BeLue R, et al. Outpatient adherence to beta-blocker therapy after acute myocardial infarction. *J Am Coll Cardiol* 2002;40:1589–95.
- Nichol MB, Venturini F, Sung JC. A critical evaluation of the methodology of the literature on medication compliance. *Ann Pharmacother* 1999;33:531–40.
- Mant J, Hicks N. Detecting differences in quality of care: the sensitivity of measures of process and outcome in treating acute myocardial infarction. *BMJ* 1995;311:793–6.

30. National Center for Health Statistics. Plan and operation of the second National Health and Nutrition Examination Survey 1976-80. Programs and collection procedures, series 1, no. 15. Hyattsville, MD: National Center for Health Statistics, 1981. (DHHS publication no. (PHS) 81-1317).
31. *Idem*. NHANES 1999-2000. (Accessed May 11, 2007, at http://www.cdc.gov/nchs/about/major/nhanes/NHANES99_00.htm.)
32. Mokdad AH, Stroup DE, Giles WH. Public health surveillance for behavioral risk factors in a changing environment: recommendations from the Behavioral Risk Factor Surveillance Team. *MMWR Recomm Rep* 2003;52(RR-9):1-12.
33. Lewington S, Clarke R, Qizilbash N, Peto R, Collins R. Age-specific relevance of usual blood pressure to vascular mortality: a meta-analysis of individual data for one million adults in 61 prospective studies. *Lancet* 2002;360:1903-13. [Erratum, *Lancet* 2003;361:1060.]
34. Critchley JA, Capewell S. Mortality risk reduction associated with smoking cessation in patients with coronary heart disease: a systematic review. *JAMA* 2003;290:86-97.
35. Law MR, Wald NJ, Thompson SG. By how much and how quickly does reduction in serum cholesterol concentration lower risk of ischaemic heart disease? *BMJ* 1994;308:367-72.
36. Briggs A, Sculpher M, Buxton M. Uncertainty in the economic evaluation of health care technologies: the role of sensitivity analysis. *Health Econ* 1994;3:95-104.
37. Design and operation of the National Hospital Discharge Survey: 1988 redesign. Programs and collection procedures, series 1, no. 39. Hyattsville, MD: National Center for Health Statistics, 2000. (DHHS publication no. (PHS) 2001-1315.)
38. The Medical Expenditure Panel Survey. Rockville, MD: Agency for Healthcare Research and Quality. (Accessed May 11, 2007, at <http://www.meps.ahrq.gov/mepsweb>.)
39. Rogers WJ, Canto JG, Lambrew CT, et al. Temporal trends in the treatment of over 1.5 million patients with myocardial infarction in the U.S. from 1990 through 1999: the National Registry of Myocardial Infarction 1, 2 and 3. *J Am Coll Cardiol* 2000;36:2056-63.
40. Antithrombotic Trialists' Collaboration. Collaborative meta-analysis of randomized trials of antiplatelet therapy for prevention of death, myocardial infarction, and stroke in high risk patients. *BMJ* 2002;324:71-86.
41. Department of Health and Human Services. International classification of diseases, 9th revision, clinical modification: ICD-9-CM. Washington, DC. Department of Health and Human Services, 1980. (DHHS publication no. (PHS) 80-1260.)
42. International statistical classification of diseases and related health problems, 10th rev., ICD-10. Geneva: World Health Organization, 1992.
43. Doliszny KM, Luepker RV, Burke GL, Pryor DB, Blackburn H. Estimated contribution of coronary artery bypass graft surgery to the decline in coronary heart disease mortality: the Minnesota Heart Survey. *J Am Coll Cardiol* 1994;24:95-103.
44. Cooper K, Davies R, Roderick P, Chase D, Raftery J. The development of a simulation model of the treatment of coronary heart disease. *Health Care Manag Sci* 2002;5:259-67.
45. Flegal KM, Graubard BI, Williamson DF, Gail MH. Excess deaths associated with underweight, overweight, and obesity. *JAMA* 2005;293:1861-7.
46. Olshansky SJ, Passaro DJ, Hershow RC, et al. A potential decline in life expectancy in the United States in the 21st century. *N Engl J Med* 2005;352:1138-45.
47. Unal B, Critchley JA, Fidan D, Capewell S. Life-years gained from modern cardiovascular treatments and population risk factor changes in England and Wales, 1981-2000. *Am J Public Health* 2005;95:103-8.
48. McAlister FA. Relative treatment effects are consistent across the spectrum of underlying risks . . . usually. *Int J Epidemiol* 2002;31:76-7.
49. Hippisley-Cox J, Coupland C. Effect of combinations of drugs on all cause mortality in patients with ischaemic heart disease: nested case-control analysis. *BMJ* 2005;330:1059-63. [Erratum, *BMJ* 2006;332:912.]

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